

Correction to Proposition 6.13 in ‘Elements of Causal Inference’

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Remark 1 *I am omitting some measure theoretic details in the arguments and only refer to the relevant literature when the statements are non-trivial. I hope that the arguments are clear enough, so that you could add the details if you wanted to. In that case, the definition of SCMs by Bongers et al. [2021] might be a good starting point.*

Definition 2 *Let \mathfrak{C} be an SCM over variables \mathbf{X} with $X, Y \in \mathbf{X}$. We say that there is a (total) causal effect from X to Y if there exists a random variable \tilde{N}_X s.t.*

$$X \not\perp\!\!\!\perp Y \quad \text{in } P^{\mathfrak{C}; do(X:=\tilde{N}_X)}.$$

Proposition 3 *Let \mathfrak{C} be an SCM over variables \mathbf{X} with $X, Y \in \mathbf{X}$. Consider*

- (i) *There is a total causal effect from X to Y .*
- (ii) *There are x^Δ and x^\square such that $P_Y^{\mathfrak{C}; do(X:=x^\Delta)} \neq P_Y^{\mathfrak{C}; do(X:=x^\square)}$.*
- (iii) *There is x^Δ such that $P_Y^{\mathfrak{C}; do(X:=x^\Delta)} \neq P_Y^{\mathfrak{C}}$.*
- (iv) *$X \not\perp\!\!\!\perp Y$ in $P_{X,Y}^{\mathfrak{C}; do(X:=\tilde{N}_X)}$ for any \tilde{N}_X whose distribution has full support.*

We then have (iv) \Rightarrow (i) \Leftrightarrow (ii) \Rightarrow (iii).

To prove this proposition, we first consider the following lemma.

Lemma 4 *Let \mathfrak{C} be an SCM over variables \mathbf{X} with $X, Y \in \mathbf{X}$ and consider a noise variable \tilde{N}_X . Then,*

$$(P^{\mathfrak{C}; do(X:=x)}(Y \in \cdot))_x \text{ is a cond. distribution for } Y \text{ given } X \text{ in } P^{\mathfrak{C}; do(X:=\tilde{N}_X)}.$$

*Many thanks to Joris Mooij and Michael Law for helpful discussions and providing ideas for Example 5(ii) and (i), respectively.

Proof. Let \mathcal{G} be the graph induced by the SCM. Substituting, iteratively, the structural assignments into each other (but not the one for X), we obtain that there is a function g s.t.

$$Y = g(X, N_{A_Y^X}),$$

where $A_Y^X := \mathbf{AN}_Y^{\mathcal{G}^*} \setminus \{X\}$ and \mathcal{G}^* equals \mathcal{G} after removing the edges incoming into X . We have that $X \perp\!\!\!\perp N_{A_Y^X}$ in $P^{\mathfrak{C};do(X:=\tilde{N}_X)}$. Therefore, the substitution theorem of conditional probabilities [Rønn-Nielsen and Hansen, 2014, Theorem 2.1.1] states that

$$(P^{\mathfrak{C};do(X:=\tilde{N}_X)}(g(x, N_{A_Y^X}) \in \cdot)_x \tag{1}$$

is a conditional probability for Y given X in $P^{\mathfrak{C};do(X:=\tilde{N}_X)}$. But (1) equals

$$(P^{\mathfrak{C};do(X:=x)}(Y \in \cdot))_x,$$

which concludes the proof of Lemma 4. \square

We are now ready to prove Proposition 3.

Proof. The implications $(iv) \Rightarrow (i)$ and $(ii) \Rightarrow (iii)$ are trivial.

To prove $(i) \Rightarrow (ii)$, assume that (ii) does not hold. Then, there exists a function c s.t. for all x and for all measurable B

$$P^{\mathfrak{C};do(X:=x)}(Y \in B) = c(B).$$

Thus, for all measurable A and B , we have

$$\begin{aligned} P^{\mathfrak{C};do(X:=\tilde{N}_X)}(X \in A, Y \in B) &= \int_A P^{\mathfrak{C};do(X:=x)}(Y \in B) dP_X^{\mathfrak{C};do(X:=\tilde{N}_X)}(x) \\ &= c(B)P^{\mathfrak{C};do(X:=\tilde{N}_X)}(X \in A), \end{aligned}$$

where for the first equality we have made use of Lemma 4. But this implies

$$c(B) = P^{\mathfrak{C};do(X:=\tilde{N}_X)}(Y \in B),$$

which shows that X and Y are independent in $P^{\mathfrak{C};do(X:=\tilde{N}_X)}$.

To prove $(ii) \Rightarrow (i)$, assume that (ii) holds and let x^Δ and x^\square be values and B be a measurable set such that

$$P_Y^{\mathfrak{C};do(X:=x^\Delta)}(B) \neq P_Y^{\mathfrak{C};do(X:=x^\square)}(B).$$

Consider a random variable \tilde{N}_X with $P(\tilde{N}_X = x^\Delta) = 1/2 = P(\tilde{N}_X = x^\square)$. We then have

$$\begin{aligned} P^{\mathfrak{C};do(X:=\tilde{N}_X)}(X = x^\Delta, Y \in B) &= \int_{\{x^\Delta\}} P^{\mathfrak{C};do(X:=x)}(Y \in B) dP_X^{\mathfrak{C};do(X:=\tilde{N}_X)}(x) \\ &= \frac{1}{2} P^{\mathfrak{C};do(X:=x^\Delta)}(Y \in B) \\ &\neq \frac{1}{2} P^{\mathfrak{C};do(X:=x^\square)}(Y \in B) \\ &= P^{\mathfrak{C};do(X:=\tilde{N}_X)}(X = x^\square, Y \in B), \end{aligned}$$

where, again, we have made use of Lemma 4. This implies that X and Y are not independent in $P^{\mathfrak{C};do(X:=\tilde{N}_X)}$ (if they were, then the first and last term should both equal $1/2 \cdot P^{\mathfrak{C};do(X:=\tilde{N}_X)}(Y \in B)$). This concludes the proof of Proposition 3. \square

Example 5 (i) Let \mathfrak{C} be an SCM over (X, Y) with

$$X := N_X \tag{2}$$

$$Y := 1_0(X) + N_Y, \tag{3}$$

where N_X and N_Y are i.i.d. $\mathcal{N}(0, 1)$ and $1_0(0) = 1$ and $1_0(x) = 0$ for all $x \neq 0$. Then, (ii) of Proposition 3 is satisfied but not (iv).

(ii) Let \mathfrak{C} be an SCM over (X, Y) with

$$Z := N_Z \tag{4}$$

$$X := Z \tag{5}$$

$$Y := \text{sg}(X) \cdot \text{sg}(Z) + N_Y, \tag{6}$$

where N_Z and N_Y are i.i.d. $\mathcal{N}(0, 1)$ and $\text{sg}(x) = 1$ for all $x \geq 0$ and $\text{sg}(x) = -1$ for all $x < 0$. Then, (iii) of Proposition 3 is satisfied but not (ii).

References

- S. Bongers, P. Forre, J. Peters, and J. M. Mooij. Foundations of structural causal models with cycles and latent variables. *Annals of Statistics*, 49(5): 2885–2915, 2021.
- A. Rønne-Nielsen and E. Hansen. Conditioning and Markov properties, 2014.